

Continuous Distributions

Continuous random variables are used to describe random phenomena in which variables take any value in some interval. Now, we will discuss various continuous distributions.

(i) Uniform Distribution

A random variable X is uniformly distributed on the interval $[a, b]$ if its pdf is given by-

$$f(x) = \begin{cases} \frac{1}{b-a}, & \text{if } a \leq x \leq b \\ 0, & \text{otherwise} \end{cases}$$

The cdf is given by -

$$F(x) = \begin{cases} 0, & \text{if } x < a \\ \frac{x-a}{b-a}, & \text{if } a \leq x < b \\ 1, & \text{if } x \geq b \end{cases}$$

The mean and variance are :

$$E(X) = \frac{a+b}{2}$$

$$\sigma^2 = \frac{(b-a)^2}{12}$$

NOTE: For Uniform distribution,

$$\begin{aligned} P(x_1 < X < x_2) &= F(x_2) - F(x_1) \\ &= \frac{x_2 - x_1}{b-a} \end{aligned}$$

Example :

A bus arrives every 20 minutes at a specified stop beginning at 6:40 am and continuing until 8:40 am. A certain passenger does not know the schedule, but arrives randomly (uniformly distributed) between 7:00 am and 7:30 am every morning. What is the probability that the passenger waits more than 5 minutes for a bus?

Solution : It is given that bus comes to the stop at every 20 min. starting from 6:40 am. Hence, bus would be coming at -

6:40 am, 7:00 am, 7:20 am, 7:40 am etc.

Passenger arrives to bus stop anytime between 7 am and 7:30 am. The passenger has to wait for more than 5 minutes means -

* he would have arrived between 7:00 am and 7:15 am. That is he arrived after the 7 am bus left.

or * he would have arrived between 7:20 am and 7:30 am. That is, he came after 7:20 am bus left.

Let X denote the no of minutes past 7:00 am that the passenger arrives. Then, we need to compute -

$$P(0 < X < 15) + P(20 < X < 30)$$

Now, X is a random variable which follows uniform distribution in the interval $(0, 30)$. So, the required probability can be presented using cdf as -

$$F(15) + F(30) - F(20)$$

Note that in this problem,

$$a=0 \quad \text{and} \quad b=30.$$

and x values are 15, 30 and 20 in cdf.
Thus,

$$\begin{aligned} F(15) + F(30) - F(20) &= \frac{15-0}{30-0} + 1 - \frac{20-0}{30-0} \\ &= \frac{1}{2} + 1 - \frac{2}{3} \\ &= \frac{3+6-4}{6} = \frac{5}{6} \end{aligned}$$

(ii) Exponential Distribution

A random variable X is said to be exponentially distributed with the parameter $\lambda > 0$, if its pdf is -

$$f(x) = \begin{cases} \lambda e^{-\lambda x}, & \text{if } x \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

The cdf is given by -

$$F(x) = \begin{cases} 0, & \text{if } x < 0 \\ 1 - e^{-\lambda x}, & \text{if } x \geq 0 \end{cases}$$

The mean and variance are:

$$E(X) = \frac{1}{\lambda}, \quad V(X) = \frac{1}{\lambda^2}$$

NOTE: Exponential distribution is used mostly in following situations :

- * To model inter-arrival time. Here λ would be arrival per hour.
- * To model service time. Here, λ may be no. of services per minute.
- * To model life time of a component that fails instantaneously. For ex. bulb. Here λ would be failure rate.

Example:

Suppose, the life of an industrial lamp (calculated in thousands of hours) is exponentially distributed with failure rate $\lambda = \frac{1}{3}$ (That is one failure for every 3000 hours, on an average). Compute (i) the probability that the life of a lamp is more than its mean life, i.e., 3000 hrs. (ii) Probability that life of lamp is between 2000 and 3000 hours.

Solution: (i) Let X be the no of hours (in thousands) denoting life of the lamp. Then, we need to compute $P(X > 3)$.

$$\begin{aligned} P(X > 3) &= 1 - P(X \leq 3) \\ &= 1 - F(3) \\ &= 1 - (1 - e^{-\lambda_3 \cdot 3}) \\ &= e^{-1} \\ &= 0.3679 \end{aligned}$$

(ii) To find the probability of life of a lamp between 2000 to 3000 hrs, we need to compute $P(2 \leq X \leq 3)$.

Now,

$$\begin{aligned} P(2 \leq X \leq 3) &= F(3) - F(2) \\ &= (1 - e^{-\lambda_3 \cdot 3}) - (1 - e^{-\lambda_3 \cdot 2}) \\ &= -e^{-1} + e^{-2/3} \\ &= -0.3679 + 0.5134 \\ &= 0.1455 \end{aligned}$$

NOTE: Exponential distribution is "memory less". To understand the meaning of this term, consider the following situation:

Let x be the life of a component. Now, the probability that the component lives for at least $s+t$ hours, given that it has already survived s hours is given by -

$$P(x > s+t | x > s) = P(x > t)$$

Example:

Consider the example of industrial lamp given in previous example. Compute the probability that the lamp will survive for another 1000 hours given that it has already survived 2500 hours.

Solution: As per the question, we need to compute -

$$\begin{aligned} P(x > 3.5 | x > 2.5) &= P(x > 1) \\ &= 1 - F(1) \\ &= 1 - (1 - e^{-\lambda \cdot 1}) \\ &= e^{-\lambda \cdot 1} = 0.7165 \end{aligned}$$

(ii) Normal Distribution

A random variable X is said to follow normal distribution with parameters μ and σ^2 , if it has the pdf -

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}(\frac{x-\mu}{\sigma})^2}, \quad -\infty < x < \infty$$

Here, $-\infty < \mu < \infty$ and $\sigma^2 > 0$.

The normal distribution is usually denoted by -

$$X \sim N(\mu, \sigma^2)$$

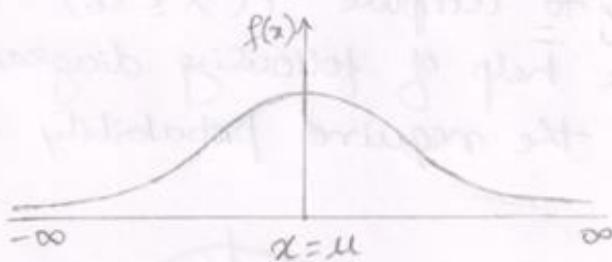
indicating X follows normal distribution. It has the following properties -

$$(i) \lim_{x \rightarrow -\infty} f(x) = 0 = \lim_{x \rightarrow \infty} f(x)$$

(ii) Normal pdf is symmetric about μ . That is -

$$f(\mu-x) = f(\mu+x)$$

The symmetry of normal distribution is shown using following diagram.



- (iii) The maximum value of pdf occurs at $x = \mu$.
- (iv) The mean, and mode are equal for normal variate.
- (v) If a random variable $X \sim N(\mu, \sigma^2)$, then

$$Z = \frac{X-\mu}{\sigma} \sim N(0, 1)$$

That is, if X is a normal variate, then $\frac{X-\mu}{\sigma}$ is a normal variate with mean as zero and variance as 1. Now, Z is called as standard normal variate. The area under standard normal curve is 1.

NOTE: For solving many problems using normal distribution, we make use of standard normal variate and the pre-defined probability table is available for standard normal variate.

Example 1.

Let X be a normal variate such that,

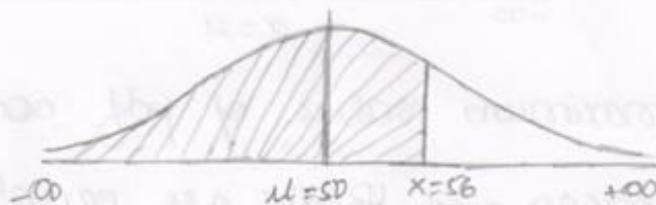
$$X \sim N(50, 9).$$

Then compute $P(X \leq 56)$.

Solution: It is given that $X \sim N(50, 9)$. That means,

$$\mu = 50 \quad \text{and} \quad \sigma^2 = 9 \\ \Rightarrow \sigma = 3$$

We need to compute $P(X \leq 56)$. This can be understood with the help of following diagram. The shaded area indicates the required probability of $X \leq 56$.



$$\text{Now, } Z = \frac{X-\mu}{\sigma} = \frac{56-50}{3} = 2 \sim N(0, 1)$$

Now, our requirement is $P(Z \leq 2)$. So, refer the standard normal distribution table (which is cumulative), and we get the value as 0.97725.
 Thus, $P(X \leq 56) = 0.97725$

Example 2:

The time required to load an ocean-going vessel, X , is distributed as $N(12, 4)$. Compute the probability that the vessel will be loaded

- (i) in less than 10 hours
- (ii) more than 12 hours
- (iii) in between 10 and 12 hours,

Solution: It is given that,

$$X \sim N(12, 4)$$

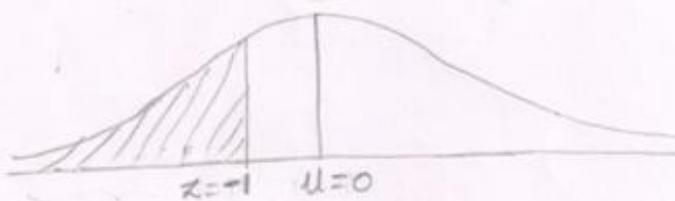
That is, $\mu = 12$ and $\sigma^2 = 4$.

$$\text{Now, } Z = \frac{X - 12}{\sqrt{4}} \sim N(0, 1)$$

(i) To compute the probability that the vessel will be loaded in less than 10 hours:

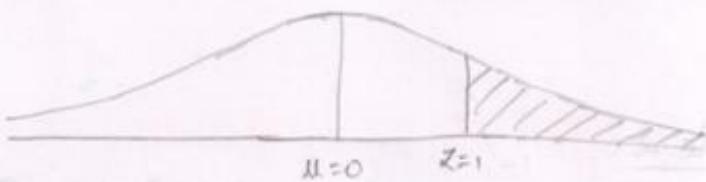
$$P(X \leq 10) \Rightarrow Z = \frac{10 - 12}{\sqrt{4}} = -1 \sim N(0, 1)$$

The standard normal curve would be—



We know that normal distribution is symmetric. So,
 $P(X \leq t) = 1 - P(X \geq t)$.

That is, in this example, we can have required curve as-



Using cumulative standard normal table, we will get-

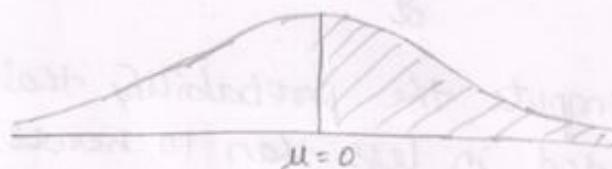
$$\begin{aligned}P(Z \leq -1) &= 1 - P(Z \geq 1) \\&= 1 - 0.84134 \\&= 0.15866\end{aligned}$$

(ii) More than 12 hours -

$$P(X \geq 12) \Rightarrow Z = \frac{12-12}{\sigma} = 0$$

i.e. we need to compute, $P(Z \geq 0)$.

The curve would be -



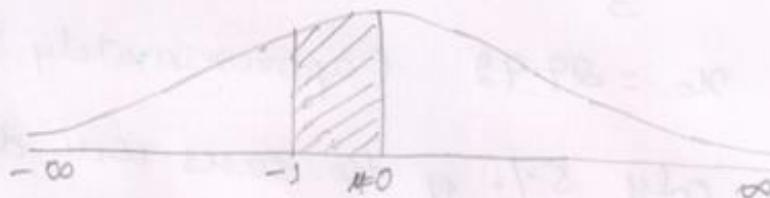
Without referring to table, just by intuition, we can say that the required probability is 0.50.

That is -

$$P(X \geq 12) = 0.5$$

(iii) Vessel is loaded between 10 and 12 hours.

We need to compute $P(10 \leq X \leq 12)$. Using standard normal variate, it would be $P(-1 \leq Z \leq 0)$. The curve is -



Now using symmetric nature of normal distribution we say that required probability will be -

$$(\text{area from } -\infty \text{ to } 0) - (\text{area from } -\infty \text{ to } -1)$$

$$= 0.5 - 0.15866$$

$$= 0.34134$$

Example 3:

Lead-time demand, X , for an item has distribution $N(25, 9)$. Compute the value of leadtime that will be exceeded only 5% of the time.

Solution: Here, we need to find some value x_0 such that $P(X > x_0) = 0.05$.

$$\begin{aligned} \text{i.e., } P(X > x_0) &= P(Z > \frac{x_0 - 25}{3}) \\ &= 1 - P\left(\frac{x_0 - 25}{3}\right) = 0.05 \end{aligned}$$

$$\Rightarrow P\left(Z < \frac{x_0 - 25}{3}\right) = 0.95.$$

Now, search the cumulative standard normal table for the value 0.95. It is found at 1.64.

$$\text{Thus } P(Z < 1.64) = 0.95$$

$$\Rightarrow \frac{x_0 - 25}{3} = 1.64$$

$$\Rightarrow x_0 = 29.92 \text{ (approximately 30)}$$

Thus, only 5% of the cases will demand delivery lead-time exceed available inventory, if the order to purchase is made when the stock reaches 30.

Gamma Distribution

Gamma distribution of a random variable X is defined using a Gamma function. A gamma function for any $\beta > 0$ is given as -

$$\Gamma(\beta) = \int_0^{\infty} x^{\beta-1} \cdot e^{-x} dx$$

Note that $\Gamma(1) = 1$ and $\Gamma(\beta) = (\beta-1)!$ when β is integer.

A random variable X is said to follow gamma distribution with parameters β and θ , if its pdf is -

$$f(x) = \begin{cases} \frac{\beta^\theta}{\Gamma(\beta)} (\beta x)^{\beta-1} e^{-\beta x}, & \text{if } x > 0 \\ 0, & \text{otherwise} \end{cases}$$

In gamma distribution, the parameter β is called as shape parameter and θ is called as scale parameter.

The mean and variance are given by -

$$E(X) = \frac{1}{\theta} \quad \text{and} \quad V(X) = \frac{1}{\beta\theta^2}$$

The cdf of X is given by -

$$F(x) = \begin{cases} 1 - \int_x^\infty \frac{\beta\theta}{\Gamma\beta} \cdot (\beta\theta t)^{\beta-1} e^{-\beta\theta t} dt, & \text{if } x > 0 \\ 0, & \text{if } x \leq 0 \end{cases}$$

Beta Distribution

A random variable X is beta distributed with parameters $\beta_1 > 0$ and $\beta_2 > 0$, if the pdf is -

$$f(x) = \begin{cases} \frac{x^{\beta_1-1} (1-x)^{\beta_2-1}}{B(\beta_1, \beta_2)}, & 0 < x < 1 \\ 0, & \text{otherwise.} \end{cases}$$

Here,

$$B(\beta_1, \beta_2) = \frac{\Gamma(\beta_1) \cdot \Gamma(\beta_2)}{\Gamma(\beta_1 + \beta_2)}$$

The pdf, $f(x)$, of beta distribution given above is defined in the range of $(0, 1)$. But, in general, distribution is defined in the

range (a, b) , where $a < b$. Then, we can define a new random variable as -

$$Y = a + (b-a)X.$$

Now, mean and variance of Y are given as -

$$E(Y) = a + (b-a)\left(\frac{\beta_1}{\beta_1 + \beta_2}\right)$$

and -

$$\text{Var}(Y) = (b-a)^2 \left(\frac{\beta_1 \beta_2}{(\beta_1 + \beta_2)^2 (\beta_1 + \beta_2 + 1)} \right)$$

Poisson Process

Consider random events such as-

- * arrival of jobs at a job shop
- * arrival of email to a mail server
- * arrival of boats to a dock
- * arrival of calls to a call center
- * breakdown of machines in a large factory etc.

These events are described by a counting function $N(t)$ defined for all $t \geq 0$. This counting function will represent the no of events that occurred in $[0, t]$. Here, the time zero indicates the time at which the observation began. For each interval $[0, t]$, the value $N(t)$ is an observation of a random variable which takes the integer values $0, 1, 2, \dots$

The counting process $\{N(t), t \geq 0\}$ is said to be a Poisson process with mean λ if the following assumptions are satisfied:

(i) Arrivals occur one at a time.

(ii) $\{N(t), t \geq 0\}$ has stationary increments: The distribution of the no of arrivals between t and $t+s$ depends only on the length of the interval s , not on the starting point t . Thus arrivals are completely at random without rush or slack periods.

(iii) $\{N(t), t \geq 0\}$ has independent increments: The no of arrivals during non-overlapping time intervals are independent random variables.

If arrivals according to a Poisson process, satisfying above three assumptions, it can be shown that the probability that $N(t)$ is equal to n is given by-

$$P[N(t) = n] = \frac{e^{-\lambda t} (\lambda t)^n}{n!}, \quad n=0,1,2, \dots \quad t \geq 0$$

Thus, $N(t)$ has the poisson distribution with parameter λt . Thus, mean and variance are -

$$E(N(t)) = \lambda t$$

$$V(N(t)) = \lambda t$$

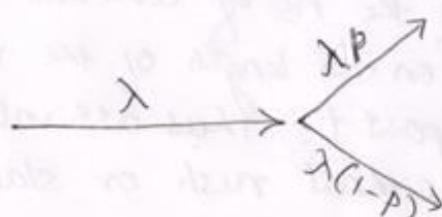
Properties of a Poisson Process

The two properties of a poisson process which are useful in discrete system simulation are -

(i) Random Splitting

(ii) Pooled Process

(i) Random Splitting: Consider a poisson process $\{N(t), t \geq 0\}$ having rate λ as shown in following figure-



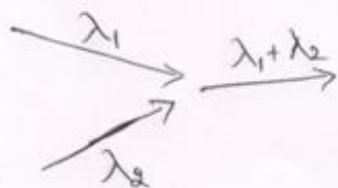
Suppose, each event is classified as type I event with probability p and type II event with probability $1-p$. Let $N_1(t)$ and $N_2(t)$ be the random variables denoting no of type I and type II events respectively.

Note that the events are occurring in $[0, t]$ and,

$$N(t) = N_1(t) + N_2(t)$$

The two events $N_1(t)$ and $N_2(t)$ are independent Poisson processes having rates λp and $\lambda(1-p)$.

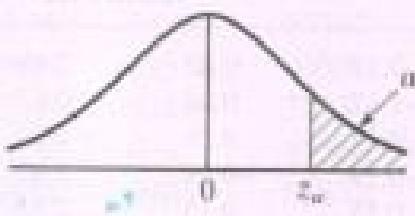
(ii) Pooled Process: Consider the opposite situation of random splitting, the pooling of two arrival streams as shown -



If $N_1(t)$ and $N_2(t)$ are two random variables representing independent Poisson processes with rates λ_1 and λ_2 , then

$N(t) = N_1(t) + N_2(t)$ is a Poisson process with rate $\lambda_1 + \lambda_2$.

Table A.3 Cumulative Normal Distribution



$$\phi(z_\alpha) = \int_{-\infty}^{z_\alpha} \frac{1}{\sqrt{2\pi}} e^{-u^2/2} du = 1 - \alpha$$

z_α	0.00	0.01	0.02	0.03	0.04	z_α
0.0	0.500 00	0.503 99	0.507 98	0.511 97	0.515 95	0.0
0.1	0.539 83	0.543 79	0.547 76	0.551 72	0.555 67	0.1
0.2	0.579 26	0.583 17	0.587 06	0.590 95	0.594 83	0.2
0.3	0.617 91	0.621 72	0.625 51	0.629 30	0.633 07	0.3
0.4	0.655 42	0.659 10	0.662 76	0.666 40	0.670 03	0.4
0.5	0.691 46	0.694 97	0.698 47	0.701 94	0.705 40	0.5
0.6	0.725 75	0.729 07	0.732 37	0.735 65	0.738 91	0.6
0.7	0.758 03	0.761 15	0.764 24	0.767 30	0.770 35	0.7
0.8	0.788 14	0.791 03	0.793 89	0.796 73	0.799 54	0.8
0.9	0.815 94	0.818 59	0.821 21	0.823 81	0.826 39	0.9
1.0	0.841 34	0.843 75	0.846 13	0.848 49	0.850 83	1.0
1.1	0.864 33	0.866 50	0.868 64	0.870 76	0.872 85	1.1
1.2	0.884 93	0.886 86	0.888 77	0.890 65	0.892 51	1.2
1.3	0.903 20	0.904 90	0.906 58	0.908 24	0.909 88	1.3
1.4	0.919 24	0.920 73	0.922 19	0.923 64	0.925 06	1.4
1.5	0.933 19	0.934 48	0.935 74	0.936 99	0.938 22	1.5
1.6	0.945 20	0.946 30	0.947 38	0.948 45	0.949 50	1.6
1.7	0.955 43	0.956 37	0.957 28	0.958 18	0.959 07	1.7
1.8	0.964 07	0.964 85	0.965 62	0.966 37	0.967 11	1.8
1.9	0.971 28	0.971 93	0.972 57	0.973 20	0.973 81	1.9
2.0	0.977 25	0.977 78	0.978 31	0.978 82	0.979 32	2.0
2.1	0.982 14	0.982 57	0.983 00	0.983 41	0.983 82	2.1
2.2	0.986 10	0.986 45	0.986 79	0.987 13	0.987 45	2.2
2.3	0.989 28	0.989 56	0.989 83	0.990 10	0.990 36	2.3
2.4	0.991 80	0.992 02	0.992 24	0.992 45	0.992 66	2.4
2.5	0.993 79	0.993 96	0.994 13	0.994 30	0.994 46	2.5
2.6	0.995 34	0.995 47	0.995 60	0.995 73	0.995 85	2.6
2.7	0.996 53	0.996 64	0.996 74	0.996 83	0.996 93	2.7
2.8	0.997 44	0.997 52	0.997 60	0.997 67	0.997 74	2.8
2.9	0.998 13	0.998 19	0.998 25	0.998 31	0.998 36	2.9
3.0	0.998 65	0.998 69	0.998 74	0.998 78	0.998 82	3.0
3.1	0.999 03	0.999 06	0.999 10	0.999 13	0.999 16	3.1
3.2	0.999 31	0.999 34	0.999 36	0.999 38	0.999 40	3.2
3.3	0.999 52	0.999 53	0.999 55	0.999 57	0.999 58	3.3
3.4	0.999 66	0.999 68	0.999 69	0.999 70	0.999 71	3.4
3.5	0.999 77	0.999 78	0.999 78	0.999 79	0.999 80	3.5
3.6	0.999 84	0.999 85	0.999 85	0.999 86	0.999 86	3.6
3.7	0.999 89	0.999 90	0.999 90	0.999 90	0.999 91	3.7
3.8	0.999 93	0.999 93	0.999 93	0.999 94	0.999 94	3.8
3.9	0.999 95	0.999 95	0.999 96	0.999 96	0.999 96	3.9

continues...

Table A.3 Continued

z_α	0.05	0.06	0.07	0.08	0.09	z_α
0.0	0.519 94	0.523 92	0.527 90	0.531 88	0.535 86	0.0
0.1	0.559 62	0.563 56	0.567 49	0.571 42	0.575 34	0.1
0.2	0.598 71	0.602 57	0.606 42	0.610 26	0.614 09	0.2
0.3	0.636 83	0.640 58	0.644 31	0.648 03	0.651 73	0.3
0.4	0.673 64	0.677 24	0.680 82	0.684 38	0.687 93	0.4
0.5	0.708 84	0.712 26	0.715 66	0.719 04	0.722 40	0.5
0.6	0.742 15	0.745 37	0.748 57	0.751 75	0.754 90	0.6
0.7	0.773 37	0.776 37	0.779 35	0.782 30	0.785 23	0.7
0.8	0.802 34	0.805 10	0.807 85	0.810 57	0.813 27	0.8
0.9	0.824 94	0.831 47	0.833 97	0.836 46	0.838 91	0.9
1.0	0.853 14	0.855 43	0.857 69	0.859 93	0.862 14	1.0
1.1	0.874 93	0.876 97	0.879 00	0.881 00	0.882 97	1.1
1.2	0.894 35	0.896 16	0.897 96	0.899 73	0.901 47	1.2
1.3	0.911 49	0.913 08	0.914 65	0.916 21	0.917 73	1.3
1.4	0.926 47	0.927 85	0.929 22	0.930 56	0.931 89	1.4
1.5	0.939 43	0.940 62	0.941 79	0.942 95	0.944 08	1.5
1.6	0.950 53	0.951 54	0.952 54	0.953 52	0.954 48	1.6
1.7	0.959 94	0.960 80	0.961 64	0.962 46	0.963 27	1.7
1.8	0.967 84	0.968 56	0.969 26	0.969 95	0.970 62	1.8
1.9	0.974 41	0.975 00	0.975 58	0.976 15	0.976 70	1.9
2.0	0.979 82	0.980 30	0.980 77	0.981 24	0.981 69	2.0
2.1	0.984 22	0.984 61	0.985 00	0.985 37	0.985 74	2.1
2.2	0.987 78	0.988 09	0.988 40	0.988 70	0.988 99	2.2
2.3	0.990 61	0.990 86	0.991 11	0.991 34	0.991 58	2.3
2.4	0.992 86	0.993 05	0.993 24	0.993 43	0.993 61	2.4
2.5	0.994 61	0.994 77	0.994 92	0.995 06	0.995 20	2.5
2.6	0.995 98	0.996 09	0.996 21	0.996 32	0.996 43	2.6
2.7	0.997 02	0.997 11	0.997 20	0.997 28	0.997 36	2.7
2.8	0.997 81	0.997 88	0.997 95	0.998 01	0.998 07	2.8
2.9	0.998 41	0.998 46	0.998 51	0.998 56	0.998 61	2.9
3.0	0.998 86	0.998 89	0.998 93	0.998 97	0.999 00	3.0
3.1	0.999 18	0.999 21	0.999 24	0.999 26	0.999 29	3.1
3.2	0.999 42	0.999 44	0.999 46	0.999 48	0.999 50	3.2
3.3	0.999 60	0.999 61	0.999 62	0.999 64	0.999 65	3.3
3.4	0.999 72	0.999 73	0.999 74	0.999 75	0.999 76	3.4
3.5	0.999 81	0.999 81	0.999 82	0.999 83	0.999 83	3.5
3.6	0.999 87	0.999 87	0.999 88	0.999 88	0.999 89	3.6
3.7	0.999 91	0.999 92	0.999 92	0.999 92	0.999 92	3.7
3.8	0.999 94	0.999 94	0.999 95	0.999 95	0.999 95	3.8
3.9	0.999 96	0.999 96	0.999 96	0.999 97	0.999 97	3.9